

# Processes of class $(\Sigma)$ , last passage times and drawdowns

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## Abstract

We propose a general framework to study last passage times, suprema and drawdowns of a large class of stochastic processes. A central role in our approach is played by processes of class  $(\Sigma)$ . After investigating convergence properties and a family of transformations that leave processes of class  $(\Sigma)$  invariant, we provide three general representation results. The first one allows to recover a process of class  $(\Sigma)$  from its final value and the last time it visited the origin. In many situations this gives access to the distribution of the last time a stochastic process hit a certain level or was equal to its running maximum. It also leads to a formula recently discovered by Madan, Roynette and Yor expressing put option prices in terms of last passage times. Our second representation result is a stochastic integral representation of certain functionals of processes of class  $(\Sigma)$ , and the third one gives a formula for their conditional expectations. From the latter one can deduce the laws of a variety of interesting random variables such as running maxima, drawdowns and maximum drawdowns of suitably stopped processes. As an application we discuss the pricing and hedging of options that depend on the running maximum of an underlying price process and are triggered when the underlying price drops to a given level or alternatively, when the drawdown or relative drawdown of the underlying price attains a given height.

**Keywords:** Processes of class  $(\Sigma)$ , last passage times, drawdowns, maximum drawdowns, relative drawdowns.

**Notation:** Let  $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$  be a filtered probability space satisfying the usual assumptions. All stochastic processes  $(X_t)$  will be indexed by  $t \in \mathbb{R}_+$ .  $\bar{X}_t$  denotes the running supremum  $\sup_{u \leq t} X_u$ . All semimartingales will be assumed to be càdlàg. Equalities and inequalities between random variables are understood in the  $\mathbb{P}$ -almost sure sense. We recall that a measurable process  $(X_t)$  is said to be of class (D) if the family of random variables  $\{|X_T|1_{\{T < \infty\}} : T \text{ a stopping time}\}$  is uniformly integrable.

## 1 Introduction

We propose a general framework to study various properties of continuous-time stochastic processes which is based on the concept of processes of class  $(\Sigma)$ . Non-negative local submartingales of class  $(\Sigma)$  were introduced by Yor [26] and have been further studied by Nikeghbali [17, 18]. They are closely

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related to relative martingales (see for instance, Azéma and Yor [4] or Azéma et al. [2]). Compared to Yor [26] and Nikeghbali [17, 18] we here work with the following larger class of stochastic processes:

**Definition 1.1** *We say a stochastic process  $(X_t)$  is of class  $(\Sigma)$  if it decomposes as  $X_t = N_t + A_t$ , where*

- (1)  $(N_t)$  is a càdlàg local martingale,
- (2)  $(A_t)$  is an adapted continuous finite variation process starting at 0,
- (3)  $\int_0^t 1_{\{X_u \neq 0\}} dA_u = 0$  for all  $t \geq 0$ .

We say  $(X_t)$  is of class  $(\Sigma D)$  if it is of class  $(\Sigma)$  and of class  $(D)$ . By  $L$  we denote the random time

$$L := \sup \{t : X_t = 0\} \quad \text{with the convention } \sup \emptyset = 0.$$

According to this definition, every local martingale is of class  $(\Sigma)$  and all uniformly integrable martingales are of class  $(\Sigma D)$ . In Lemma 2.1 below it is shown that if a process  $(X_t)$  of class  $(\Sigma)$  is non-negative, then  $(A_t)$  has to be increasing and  $(X_t)$  is a local submartingale. This case includes the absolute value  $(|M_t|)$  of continuous local martingales  $(M_t)$  as well as drawdown processes  $(\overline{M}_t - M_t)$  of local martingales whose running suprema  $(\overline{M}_t)$  are continuous. It will also follow from Lemma 2.1 that for every constant  $K \in \mathbb{R}$ , the process  $(K - M_t)^+$  is of class  $(\Sigma)$  if  $(M_t)$  is a local martingale with no positive jumps. Many other processes, such as suitably transformed diffusions or the Azéma submartingale in the filtration generated by the Brownian zeros, fall into the class  $(\Sigma)$ . Lemma 2.2 below shows that the product of processes of class  $(\Sigma)$  with vanishing quadratic covariation is again of class  $(\Sigma)$ , and in Lemma 2.3 it is proved that the class  $(\Sigma)$  is stable under transformations of the form  $(X_t) \mapsto (f(A_t)X_t)$ , where  $f : \mathbb{R} \rightarrow \mathbb{R}$  is a locally bounded Borel function.

We show how to use Itô's formula and martingale techniques to derive general results for processes of class  $(\Sigma)$  with interesting consequences for a wide range of stochastic processes and related random times. In particular, our arguments do not need Markov or scaling properties. In Section 2 we provide preliminary results on the convergence, positive parts and products of processes of class  $(\Sigma)$ . Then we extend two results on processes of the form  $f(A_t)X_t$  from Nikeghbali [18] to our setup. In Section 3 we prove three general representation results. The first gives conditions under which a process  $(X_t)$  of class  $(\Sigma)$  converges to a limit  $X_\infty$  almost everywhere on the set  $\{L < \infty\}$  and can be written as

$$X_t = \mathbb{E} [X_\infty 1_{\{L \leq t\}} \mid \mathcal{F}_t], \quad t \geq 0. \tag{1.1}$$

Similarly to the case of uniformly integrable martingales, which are always of the form  $M_t = \mathbb{E} [M_\infty \mid \mathcal{F}_t]$ , formula (1.1) represents  $(X_t)$  in terms of its final value and the last time it visited zero. If  $X_\infty = 1$ , one can take expectations on both sides of (1.1) to obtain  $\mathbb{E} [X_t] = \mathbb{P}[L \leq t]$ . In situations where  $\mathbb{E} [X_t]$  can be calculated, this give access to the distribution of the random time  $L$ ; see Nikeghbali and Platen [19] for examples. If  $(M_t)$  is a non-negative local martingale without positive jumps such that  $M_t \rightarrow 0$  almost surely, then for all  $K \in \mathbb{R}_+$ , the submartingale  $(K - M_t)^+$  is of class  $(\Sigma)$  with last zero  $g_K = \sup \{t : M_t \geq K\}$ . So as a special case of (1.1), one obtains the following formula of Madan et al. [13]:

$$\mathbb{E} [(K - M_t)^+] = K \mathbb{P}[g_K \leq t], \quad t \geq 0. \tag{1.2}$$

Our second representation result shows how to write random variables of the form  $h(A_\infty)$  as stochastic integrals with respect to  $(N_t)$ , and the third one gives a formula for  $\mathbb{E} [h(A_\infty) \mid \mathcal{F}_T]$ , where  $T$  is an

arbitrary stopping time. This leads to closed form expressions for the conditional distributions of  $A_\infty$ . In the first part of Section 4 we apply the results of Section 3 to study drawdowns  $DD_t = \overline{M}_t - M_t$  and relative drawdowns  $rDD_t = 1 - M_t/\overline{M}_t$  of local martingales  $(M_t)$ . We provide formulas for the distributions of  $\overline{M}_{T_\lambda}$  for stopping times of the form  $T_\lambda = \inf \{t : M_t = \lambda(\overline{M}_t)\}$  and for the random times  $g_\lambda = \sup \{t \leq T_\lambda : M_t = \overline{M}_t\}$ , where  $\lambda$  is a Borel function. We are also able to calculate the distributions of the maximum drawdown  $\sup_{T \leq t \leq T_K} DD_t$  and maximum relative drawdown  $\sup_{T \leq t \leq T_K} rDD_t$  for arbitrary stopping times  $T$  and hitting times of the form  $T_K = \inf \{t \geq T : Y_t = K\}$ . In the second part of Section 4 we extend these results to processes  $(Y_t)$  which admit a continuous increasing function  $s$  such that  $s(Y_t)$  is a local martingale. In Section 5 we discuss applications in financial modelling and risk management. First we discuss the pricing and hedging of options that depend on the running maximum of an underlying price process  $(M_t)$  and are triggered when  $(M_t)$  drops to a prespecified level  $c \in [0, M_0)$  or when  $DD_t$  or  $rDD_t$  reach a value  $c > 0$ . Then we calculate distributions of various random variables related to diffusions of the form  $dY_t = \mu(Y_t)dt + \sigma(Y_t)dB_t$ . An earlier result of Lehoczky [12] appears as a special case.

## 2 Preliminaries

We start by studying positive and negative parts of processes of class  $(\Sigma)$ , non-negative processes of class  $(\Sigma)$  and the convergence of  $X_t$  for  $t \rightarrow \infty$ .

**Lemma 2.1** *Let  $(X_t)$  be a process of class  $(\Sigma)$ . Then*

- (1)  $(X_t^+)$  and  $(X_t^-)$  are local submartingales.
- (2) If  $(X_t)$  has no negative jumps, then  $(X_t^+)$  is again of class  $(\Sigma)$ . If  $(X_t)$  has no positive jumps, then  $(X_t^-)$  is of class  $(\Sigma)$ .
- (3) If  $(X_t)$  is non-negative, then it is a local submartingale with  $A_t = \sup_{u \leq t} (-N_u) \vee 0$ .
- (4) If  $(X_t)$  is of class  $(\Sigma D)$ , then  $(N_t)$  is a uniformly integrable martingale and  $(A_t)$  of integrable total variation; in particular, there exist integrable random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty$ ,  $N_t \rightarrow N_\infty$ ,  $A_t \rightarrow A_\infty$  almost surely and in  $L^1$ .

*Proof.* (1) Since  $(A_t)$  is continuous, one has  $\int_0^t 1_{\{X_{u-} > 0\}} dA_u = \int_0^t 1_{\{X_u > 0\}} dA_u = 0$ . So Tanaka's formula yields

$$X_t^+ = X_0^+ + \int_0^t 1_{\{X_{u-} > 0\}} dX_u + V_t = X_0^+ + \int_0^t 1_{\{X_{u-} > 0\}} dN_u + V_t \quad (2.1)$$

for the increasing finite variation process

$$V_t = \sum_{0 < u \leq t} 1_{\{X_{u-} \leq 0\}} X_u^+ + \sum_{0 < u \leq t} 1_{\{X_{u-} > 0\}} X_u^- + \frac{1}{2} l_t$$

and  $(l_t)$  the local time of  $(X_t)$  at 0 (see, for instance, Protter [24]). This shows that  $(X_t^+)$  is a local submartingale. The same is true for  $(X_t^-)$  because  $(-X_t)$  is also of class  $(\Sigma)$ .

(2) If  $(X_t)$  has no negative jumps, equation (2.1) reduces to

$$X_t^+ = X_0^+ + \int_0^t 1_{\{X_{u-} > 0\}} dN_u + \sum_{0 < u \leq t} 1_{\{X_{u-} \leq 0\}} X_u^+ + \frac{1}{2} l_t. \quad (2.2)$$

$(\int_0^t 1_{\{X_{u-} > 0\}} dN_u)$  is a local martingale and the local time  $(l_t)$  is continuous and has the property  $\int_0^t 1_{\{X_u \neq 0\}} dl_u = 0$  for all  $t \geq 0$ . It remains to show that the process  $Y_t = \sum_{0 < u \leq t} 1_{\{X_{u-} \leq 0\}} X_u^+$  can be decomposed into the sum of a local martingale and an adapted continuous increasing process  $(C_t)$  with  $C_0 = 0$  and

$$\int_0^t 1_{\{X_u^+ \neq 0\}} dC_u = 0 \quad \text{for all } t \geq 0. \quad (2.3)$$

Since  $(N_t)$  and  $(\int_0^t 1_{\{X_{u-} > 0\}} dN_u)$  are local martingales and  $(A_t)$  is continuous, there exists a sequence of stopping times  $(T_n)_{n \in \mathbb{N}}$  increasing to  $\infty$  such that

$$\mathbb{E} [(X_{T_n})^+] = \mathbb{E} [(N_{T_n} + A_{T_n})^+] < \infty \quad \text{and} \quad \mathbb{E} \left[ \int_0^{T_n} 1_{\{X_{u-} > 0\}} dN_u \right] = 0$$

for all  $n \in \mathbb{N}$ . So it follows from (2.2) that  $\mathbb{E} [Y_{T_n}] \leq \mathbb{E} [(X_{T_n})^+] < \infty$  for all  $n \in \mathbb{N}$ . Hence, by Theorem VI.80 of Dellacherie and Meyer [7], there exists a right-continuous increasing predictable process  $(C_t)$  starting at 0 such that  $Y_t - C_t$  is a local martingale. Since  $(A_t)$  is continuous, the jumps of  $(X_t)$  coincide with those of  $(N_t)$ . Due to the local martingale property of  $(N_t)$  and the fact that the jumps are positive, they have to occur at totally inaccessible stopping times. From Theorem VI.76 of Dellacherie and Meyer [7], one obtains  $\mathbb{E} [\Delta C_T] = \mathbb{E} [\Delta Y_T] = 0$  for every predictable stopping time  $T < \infty$ . This shows that  $(C_t)$  is continuous. Moreover, there exists a sequence of stopping times  $(R_n)_{n \in \mathbb{N}}$  increasing to  $\infty$  such that

$$\mathbb{E} \left[ \int_0^{t \wedge R_n} 1_{\{X_{u-}^+ \neq 0\}} dC_u \right] = \mathbb{E} \left[ \int_0^{t \wedge R_n} 1_{\{X_{u-}^+ \neq 0\}} dY_u \right] = \mathbb{E} \left[ \sum_{0 < u \leq t \wedge R_n} 1_{\{X_{u-}^+ \neq 0\}} 1_{\{X_{u-} \leq 0\}} X_u^+ \right] = 0$$

for all  $n \in \mathbb{N}$ . By monotone convergence one obtains

$$\mathbb{E} \left[ \int_0^t 1_{\{X_{u-}^+ \neq 0\}} dC_u \right] = \mathbb{E} \left[ \int_0^t 1_{\{X_{u-}^+ \neq 0\}} dY_u \right] = \mathbb{E} \left[ \sum_{0 < u \leq t} 1_{\{X_{u-}^+ \neq 0\}} 1_{\{X_{u-} \leq 0\}} X_u^+ \right] = 0.$$

This shows (2.3) and proves that  $(X_t^+)$  is of class  $(\Sigma)$ . That  $(X_t^-)$  is of class  $(\Sigma)$  if  $(X_t)$  has no positive jumps follows from the same arguments applied to  $(-X_t)$ .

(3) If  $(X_t)$  is non-negative, it follows from (1) that it is a local submartingale. Hence,  $A_t \geq A_u \geq -N_u \vee 0$  for all  $t \geq u$ , and therefore,  $A_t \geq \sup_{u \leq t} (-N_u) \vee 0$ . Now assume

$$\mathbb{P}[A_t > \sup_{u \leq t} (-N_u) \vee 0] > 0 \quad (2.4)$$

and introduce the random time  $T = \sup \{s \leq t : A_s = \sup_{u \leq s} (-N_u) \vee 0\}$ . Since  $(A_s)$  is continuous and  $\sup_{u \leq s} (-N_u) \vee 0$  increasing, one has  $A_T = \sup_{u \leq T} (-N_u) \vee 0$ . Moreover, since  $X_u > 0$  on the stochastic

interval  $\{(u, \omega) : T(\omega) < u \leq t\}$ , it follows from  $\int_0^t 1_{\{X_u \neq 0\}} dA_u = 0$  that  $A_t = A_T$ , a contradiction to (2.4). Hence,  $A_t = \sup_{u \leq t} (-N_u) \vee 0$ .

(4) If  $(X_t)$  is of class  $(\Sigma D)$ , then  $(X_t^+)$  and  $(X_t^-)$  are submartingales of class (D). Therefore both have a Doob–Meyer decomposition into the sum of a uniformly integrable martingale and a predictable increasing process of integrable total variation:

$$X_t^+ = N_t^1 + V_t^1, \quad X_t^- = N_t^2 + V_t^2.$$

Since the predictable finite variation part of a special semimartingale is unique, one has  $N_t = N_t^1 - N_t^2$  and  $A_t = V_t^1 - V_t^2$ . So  $(N_t)$  is a uniformly integrable martingale and  $(A_t)$  of integrable total variation. It follows that there exist integrable random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty, N_t \rightarrow N_\infty, A_t \rightarrow A_\infty$  almost surely and in  $L^1$ .  $\square$

The next lemma shows that the product of processes of class  $(\Sigma)$  with vanishing quadratic covariations is again of class  $(\Sigma)$ .

**Lemma 2.2** *Let  $(X_t^1), \dots, (X_t^n)$  be processes of class  $(\Sigma)$  such that  $[X^i, X^j]_t = 0$  for  $i \neq j$ . Then  $\prod_{i=1}^n X_t^i$  is again of class  $(\Sigma)$ .*

*Proof.* Since  $[X^1, X^2]_t = 0$ , integration by parts yields

$$X_t^1 X_t^2 = X_0^1 X_0^2 + \int_0^t X_{u-}^1 dN_u^2 + \int_0^t X_{u-}^2 dN_u^1 + \int_0^t X_u^1 dA_u^2 + \int_0^t X_u^2 dA_u^1.$$

$\int_0^t X_{u-}^1 dN_u^2 + \int_0^t X_{u-}^2 dN_u^1$  is a local martingale and  $\int_0^t X_u^1 dA_u^2 + \int_0^t X_u^2 dA_u^1$  a continuous finite variation process starting at 0 which only moves when  $X_t^1 = 0$  or  $X_t^2 = 0$ . Hence,  $X_t^1 X_t^2$  is of class  $(\Sigma)$ . If  $n \geq 3$ , then  $[X^1 X^2, X^3]_t = 0$ , and the lemma follows by induction.  $\square$

In the following lemma and the subsequent corollary we extend results of Nikeghbali [18] to our framework that will be needed later in the paper.

**Lemma 2.3** *Let  $(X_t)$  be a process of class  $(\Sigma)$  and  $f : \mathbb{R} \rightarrow \mathbb{R}$  a locally bounded Borel function. Denote  $F(x) = \int_0^x f(y) dy$ . Then the following hold:*

(1) *The process  $f(A_t)X_t$  is again of class  $(\Sigma)$  with decomposition*

$$f(A_t)X_t = f(0)X_0 + \int_0^t f(A_u) dN_u + F(A_t). \quad (2.5)$$

(2) *If  $(f(A_t)X_t)$  is of class (D), then  $M_t = f(A_t)X_t - F(A_t)$  is a uniformly integrable martingale, and therefore,*

$$f(A_T)X_T - F(A_T) = \mathbb{E}[M_\infty \mid \mathcal{F}_T] \quad \text{for every stopping time } T. \quad (2.6)$$

*Proof.* (1) It can easily be checked that  $f(A_t)X_t$  is càdlàg. To show that it is of class  $(\Sigma)$  with decomposition (2.5) we first assume that  $f$  is  $C^1$ . Then

$$f(A_t)X_t = f(0)X_0 + \int_0^t f(A_u)(dN_u + dA_u) + \int_0^t X_u f'(A_u) dA_u.$$

Since  $(X_t)$  is of class  $(\Sigma)$ , the last integral vanishes. So

$$f(A_t)X_t = f(0)X_0 + \int_0^t f(A_u)dN_u + F(A_t).$$

$f(0)X_0 + \int_0^t f(A_u)dN_u$  is a local martingale and  $F(A_t)$  a continuous adapted finite variation process starting at 0. Moreover,

$$\int_0^t 1_{\{X_u \neq 0\}} dF(A_u) = \int_0^t 1_{\{X_u \neq 0\}} f(A_u) dA_u = 0,$$

and therefore also,

$$\int_0^t 1_{\{f(A_u)X_u \neq 0\}} dF(A_u) = 0,$$

which shows that  $f(A_t)X_t$  is of class  $(\Sigma)$ . The case where  $f$  is a bounded Borel function now follows from a monotone class argument. From there it can be extended to locally bounded Borel functions by localization with a sequence of stopping times.

(2) If  $f(A_t)X_t$  is of class  $(\Sigma D)$ , it follows from Lemma 2.1 that its local martingale part  $M_t = f(A_t)X_t - F(A_t)$  is a uniformly integrable martingale. Formula (2.6) is then a consequence of Doob's optional stopping theorem.  $\square$

Note that if  $(M_t)$  is a local martingale starting at  $m \in \mathbb{R}$  such that  $(\overline{M}_t)$  is continuous, then  $X_t = \overline{M}_t - M_t$  is of class  $(\Sigma)$  with decomposition  $X_t = (m - M_t) + (\overline{M}_t - m)$ . So one obtains from Lemma 2.3 that for every locally bounded Borel function  $f : \mathbb{R} \rightarrow \mathbb{R}$  and  $F(x) = \int_0^x f(y)dy$ , the process

$$F(A_t) - f(A_t)X_t = F(\overline{M}_t - m) - f(\overline{M}_t - m)(\overline{M}_t - M_t)$$

is again a local martingale. This transformation was used by Azéma and Yor [3] in their solution of the Skorokhod embedding problem. In Carraro et al. [6] it is studied for max-continuous semimartingales.

One can use Lemma 2.3 to calculate the probability that processes of the form  $f(A_t)X_t$  stay below a given constant, which without loss of generality, can be taken to be 1. This will prove useful in the study of drawdowns and relative drawdowns in Section 4.

**Corollary 2.4** *Let  $(X_t)$  be a non-negative process of class  $(\Sigma)$  with no positive jumps such that  $A_\infty = \infty$ ,  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  a Borel function and  $T < \infty$  a stopping time. Then*

$$\begin{aligned} \mathbb{P}[f(A_t)X_t < 1 \text{ for all } t \geq T \mid \mathcal{F}_T] &= \mathbb{P}[f(A_t)X_t \leq 1 \text{ for all } t \geq T \mid \mathcal{F}_T] \\ &= (1 - f(A_T)X_T)^+ \exp\left(-\int_{A_T}^\infty f(x)dx\right). \end{aligned} \tag{2.7}$$

Moreover, in both of the following two cases:

- (1)  $K$  is an  $\mathcal{F}_T$ -measurable random variable such that  $K > A_T$  and  $T_K = \inf\{t : A_t \geq K\}$
- (2)  $K$  is an  $\mathcal{F}_T$ -measurable random variable such that  $K \geq A_T$  and  $T_K = \inf\{t : A_t > K\}$ ,

one has

$$\begin{aligned} \mathbb{P}[f(A_t)X_t < 1 \text{ for all } t \in [T, T_K] \mid \mathcal{F}_T] &= \mathbb{P}[f(A_t)X_t \leq 1 \text{ for all } t \in [T, T_K] \mid \mathcal{F}_T] \\ &= (1 - f(A_T)X_T)^+ \exp\left(-\int_{A_T}^K f(x)dx\right). \end{aligned} \tag{2.8}$$

*Proof.* Let us first assume that  $f$  is bounded and  $F(\infty) = \int_0^\infty f(y)dy < \infty$ . Then one obtains from part (1) of Lemma 2.3 that  $Y_t = f(A_t)X_t$  is a non-negative process of class  $(\Sigma)$  with no positive jumps. For a given stopping time  $T < \infty$ , denote  $R = \inf \{t \geq T : Y_t \geq 1\}$ . By (2.5),  $Y_t$  decomposes as

$$Y_t = f(0)X_0 + \int_0^t f(A_u)dN_u + F(A_t)$$

and

$$e^{F(A_t)}Y_t = f(0)X_0 + \int_0^t e^{F(A_u)}f(A_u)dN_u + e^{F(A_t)}.$$

is again of class  $(\Sigma)$ . In particular,  $e^{F(A_t)}(1 - Y_t)$  is a local martingale and

$$M_t = 1_{\{t \geq T\}} \left( e^{F(A_{R \wedge t})}(1 - Y_{t \wedge R}) - e^{F(A_T)}(1 - Y_T) \right)$$

a bounded martingale such that  $M_0 = 0$  and  $M_t \rightarrow M_\infty$  almost surely and in  $L^1$  for an integrable random variable  $M_\infty$ . Note that  $M_\infty = 0$  on  $\{T = R\}$  and  $M_\infty = -e^{F(A_T)}(1 - Y_T)$  on  $\{T < R < \infty\}$ . Moreover, since  $A_\infty 1_{\{L < \infty\}} = A_L 1_{\{L < \infty\}}$  is real-valued, it follows from  $A_\infty = \infty$  that  $L = \infty$ . Hence, there exists a sequence  $(T_n)_{n \in \mathbb{N}}$  of stopping times that increase to  $\infty$  almost surely such that  $Y_{T_n} = 0$  for all  $n \in \mathbb{N}$ , and one obtains

$$M_\infty = \lim_{n \rightarrow \infty} e^{F(A_{T_n})}(1 - Y_{T_n}) - e^{F(A_T)}(1 - Y_T) = e^{F(\infty)} - e^{F(A_T)}(1 - Y_T)$$

almost everywhere on  $\{R = \infty\}$ . So  $\mathbb{E}[M_\infty | \mathcal{F}_T] = 0$  yields

$$e^{F(A_T)}(1 - Y_T)^+ = \mathbb{P}[R = \infty | \mathcal{F}_T]e^{F(\infty)},$$

which is equivalent to

$$\mathbb{P}[f(A_t)X_t < 1 \text{ for all } t \geq T | \mathcal{F}_T] = (1 - f(A_T)X_T)^+ \exp\left(-\int_{A_T}^\infty f(x)dx\right). \quad (2.9)$$

The equality

$$\mathbb{P}[f(A_t)X_t \leq 1 \text{ for all } t \geq T | \mathcal{F}_T] = (1 - f(A_T)X_T)^+ \exp\left(-\int_{A_T}^\infty f(x)dx\right) \quad (2.10)$$

follows from the same argument applied to the stopping time  $\tilde{R} = \inf \{t \geq T : Y_t > 1\}$ . That (2.9) and (2.10) still hold for general Borel functions  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  can be seen by approximating  $f$  with  $f^n = f \wedge n 1_{[0, n]}$ ,  $n \in \mathbb{N}$ . Note that the functions  $f^n$  increase to  $f$  and for every  $x \geq 0$  there exists an  $n_0 \in \mathbb{N}$  such that  $f^n(x) = f(x)$  for all  $n \geq n_0$ . Therefore, one has

$$\bigcap_{n \in \mathbb{N}} \{f^n(A_t)X_t < 1 \text{ for all } t \geq T\} = \{f(A_t)X_t < 1 \text{ for all } t \geq T\}$$

as well as

$$\bigcap_{n \in \mathbb{N}} \{f^n(A_t)X_t \leq 1 \text{ for all } t \geq T\} = \{f(A_t)X_t \leq 1 \text{ for all } t \geq T\}.$$

In case (2) one obtains (2.8) from (2.7) simply by setting  $f$  equal to 0 on  $(K, \infty)$ . In case (1), setting  $f$  equal to 0 on  $[K, \infty)$  gives

$$\begin{aligned} & \mathbb{P}[f(A_t)X_t < 1 \text{ for all } t \in [T, T_K) \mid \mathcal{F}_T] = \mathbb{P}[f(A_t)X_t \leq 1 \text{ for all } t \in [T, T_K) \mid \mathcal{F}_T] \\ & = (1 - f(A_T)X_T)^+ \exp\left(-\int_{A_T}^K f(x)dx\right). \end{aligned}$$

But this is equivalent to (2.8) since  $X_{T_K} = 0$ . □

### 3 Representation results

#### 3.1 Representations in terms of last passage times

The results in this subsection are inspired by a representation formula for relative martingales by Azéma and Yor [4] and the recent formula (1.2) of Madan et al. [13]. In fact, in the special case  $f \equiv 1$ , part (1) of Theorem 3.1 below follows from the proof of Proposition 2.2.a) in Azéma and Yor [4]. In part (2) of Theorem 3.1 and Corollary 3.2 we relax the integrability conditions on  $(X_t)$ . This leads to formulas involving conditional expectations of random variables which are conditionally integrable but not necessarily integrable. To cover this case, we define the conditional expectation of any random variable  $X$  with respect to a sub- $\sigma$ -algebra  $\mathcal{G}$  of  $\mathcal{F}$  by

$$\mathbb{E}[X \mid \mathcal{G}] = \sup_{m \in \mathbb{Z}} \inf_{n \in \mathbb{Z}} \mathbb{E}[m \wedge (n \vee X) \mid \mathcal{G}]. \quad (3.1)$$

Then

$$\mathbb{E}[XY \mid \mathcal{G}] = X\mathbb{E}[Y \mid \mathcal{G}]$$

for all  $\mathcal{G}$ -measurable random variables  $X$  and integrable random variables  $Y$ .

**Theorem 3.1** *Let  $(X_t)$  be a process of class  $(\Sigma)$  and  $f : \mathbb{R} \rightarrow \mathbb{R}$  a Borel function. Then the following hold:*

- (1) *If  $(X_t)$  is of class (D), then there exist integrable random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty, N_t \rightarrow N_\infty, A_t \rightarrow A_\infty$  almost surely as well as in  $L^1$  and*

$$f(A_T)X_T = \mathbb{E}[f(A_\infty)X_\infty 1_{\{L \leq T\}} \mid \mathcal{F}_T] \quad \text{for every stopping time } T. \quad (3.2)$$

- (2) *If  $q : \mathbb{R} \rightarrow \mathbb{R} \setminus \{0\}$  is a Borel function such that  $q(A_t)X_t$  is of class (D), then there exist random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty, N_t \rightarrow N_\infty, A_t \rightarrow A_\infty$  almost everywhere on  $\{L < \infty\}$  and*

$$f(A_T)X_T = \mathbb{E}[f(A_\infty)X_\infty 1_{\{L \leq T\}} \mid \mathcal{F}_T] \quad \text{for all stopping times } T < \infty. \quad (3.3)$$

*In particular, in both cases one has*

$$X_T = \mathbb{E}[X_\infty 1_{\{L \leq T\}} \mid \mathcal{F}_T] \quad \text{for all stopping times } T < \infty. \quad (3.4)$$



*Proof.* (1) If  $(X_t)$  is of class  $(\Sigma D)$ , it follows from Lemma 2.1 that  $(N_t)$  is a uniformly integrable martingale and  $(A_t)$  of integrable total variation. So there exist integrable random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty, N_t \rightarrow N_\infty, A_t \rightarrow A_\infty$  almost surely as well as in  $L^1$ , and for every stopping time  $T$ , the following trick from the proof of Proposition 2.2 of Azéma and Yor [4] can be applied: Denote

$$d_T = \inf \{t > T : X_t = 0\} \quad \text{with the convention } \inf \emptyset = \infty.$$

Since  $X_\infty 1_{\{L \leq T\}} = X_{d_T}$  and  $A_T = A_{d_T}$ , it follows from Doob's optional stopping theorem that

$$\mathbb{E} [X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] = \mathbb{E} [N_{d_T} + A_{d_T} | \mathcal{F}_T] = N_T + A_T = X_T.$$

Moreover, one has  $A_\infty = A_T$  almost everywhere on  $\{L \leq T\}$ , and therefore,

$$\mathbb{E} [f(A_\infty) X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] = \mathbb{E} [f(A_T) X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] = f(A_T) X_T.$$

(2) If there exists a Borel function  $q : \mathbb{R} \rightarrow \mathbb{R} \setminus \{0\}$  such that  $q(A_t) X_t$  is of class  $(D)$ , then  $h(x) = |q(x)| \wedge 1$  is a bounded Borel function and  $Y_t = h(A_t) X_t$  is still of class  $(D)$ . It follows from Lemma 2.3 that  $(Y_t)$  is of class  $(\Sigma D)$ . By (1),  $Y_t \rightarrow Y_\infty$  almost surely as well as in  $L^1$  and

$$Y_T = \mathbb{E} [Y_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] \quad \text{for every stopping time } T.$$

Since  $\int_0^t 1_{\{X_u \neq 0\}} dA_u = 0$  for all  $t \geq 0$ ,  $A_t$  converges to  $A_L$  almost everywhere on  $\{L < \infty\}$ . Hence, it follows from  $h \neq 0$  that  $X_t \rightarrow X_\infty = Y_\infty / h(A_L)$  and  $N_t \rightarrow N_\infty = X_\infty - A_L$  almost everywhere on  $\{L < \infty\}$ . On  $\{L = \infty\}$ , set  $X_\infty = N_\infty = A_\infty = 0$ . If  $T$  is a stopping time satisfying  $T < \infty$ , then

$$f(A_T) X_T = \frac{f(A_T)}{h(A_T)} Y_T = \mathbb{E} \left[ \frac{f(A_T)}{h(A_T)} Y_\infty 1_{\{L \leq T\}} | \mathcal{F}_T \right] = \mathbb{E} [f(A_\infty) X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T].$$

□

**Corollary 3.2** *Let  $(X_t)$  be a process of class  $(\Sigma)$  and  $f : \mathbb{R} \rightarrow \mathbb{R}$  a Borel function. Assume that at least one of the following two conditions holds:*

(1)  $(N_t)$  is a uniformly integrable martingale

(2)  $(X_t^-)$  and  $(N_t^+)$  are of class  $(D)$ .

*Then there exist random variables  $X_\infty, N_\infty, A_\infty$  such that  $X_t \rightarrow X_\infty, N_t \rightarrow N_\infty, A_t \rightarrow A_\infty$  almost everywhere on  $\{L < \infty\}$  and*

$$f(A_T) X_T = \mathbb{E} [f(A_\infty) X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] \quad \text{for all stopping times } T < \infty.$$

*In particular,*

$$X_T = \mathbb{E} [X_\infty 1_{\{L \leq T\}} | \mathcal{F}_T] \quad \text{for all stopping times } T < \infty.$$

*Proof.* In both cases  $e^{-|A_t|} X_t$  is of class  $(D)$ . So the corollary follows from part (2) of Theorem 3.1. □

**Remark 3.3** For representations of the form (3.2), (3.3) or (3.4) to hold it is not sufficient that a process  $(X_t)$  of class  $(\Sigma)$  has an almost sure finite limit  $\lim_{t \rightarrow \infty} X_t$ . For example,  $X_t = 1 - \exp(B_t - t/2)$  is of class  $(\Sigma)$  with  $X_0 = 0$  and  $\lim_{t \rightarrow \infty} X_t = 1$  almost surely. But  $X_t = \mathbb{P}[L \leq t \mid \mathcal{F}_t]$  cannot hold since there is a positive probability that  $X_t$  is negative and  $\mathbb{P}[L \leq t \mid \mathcal{F}_t]$  is always between 0 and 1.

Processes  $(X_t)$  of class  $(\Sigma)$  that are not of class (D) but satisfy (3.3) and (3.4) can be constructed from strict local martingales as follows: Take a non-negative continuous strict local martingale  $(M_t)$  starting at  $m \in \mathbb{R}_+ \setminus \{0\}$  such that  $\lim_{t \rightarrow \infty} M_t = 0$  almost surely (for instance,  $M_t = \|B_t\|_2^{-1}$  for a 3-dimensional Brownian motion starting from a point  $x \in \mathbb{R}^3 \setminus \{0\}$  and  $\|\cdot\|_2$  the Euclidean norm on  $\mathbb{R}^3$ ).  $(M_t)$  is a supermartingale but not a martingale. So there exists  $u \in \mathbb{R}_+$  such that  $\mathbb{E}[M_u] < m$ , and it follows from Lemma 2.1 and Proposition 2.3 of Elworthy et al. [10] that  $\mathbb{E}[\overline{M}_t] = \infty$  for all  $t \geq u$ . Hence,  $X_t = \overline{M}_t - M_t$  is a non-negative process of class  $(\Sigma)$  with  $\lim_{t \rightarrow \infty} X_t = \overline{M}_\infty$  almost surely and  $\mathbb{E}[X_t] = \infty$  for all  $t \geq u$ . Clearly,  $(X_t)$  satisfies condition (2) of Corollary 3.2. So

$$f(A_T)X_T = \mathbb{E}[f(A_\infty)X_\infty 1_{\{L \leq T\}} \mid \mathcal{F}_T]$$

for every Borel function  $f : \mathbb{R} \rightarrow \mathbb{R}$  and stopping time  $T < \infty$ , even though  $X_\infty$  is not integrable and the conditional expectation has to be understood in the sense of (3.1).

As a consequence of Lemma 2.1 and Theorem 3.1, one obtains the following

**Corollary 3.4** (Madan–Roynette–Yor [13])

Let  $K$  be a constant and  $(M_t)$  a local martingale with no positive jumps such that  $(M_t^-)$  is of class (D). Denote  $g_K = \sup\{t : M_t \geq K\}$ . Then

$$(K - M_T)^+ = \mathbb{E}[(K - M_\infty)^+ 1_{\{g_K \leq T\}} \mid \mathcal{F}_T], \quad (3.5)$$

for every stopping time  $T$ . In particular, if  $M_\infty = m \in \mathbb{R}$ , then

$$(K - M_T)^+ = (K - m)^+ \mathbb{P}[g_K \leq T \mid \mathcal{F}_T].$$

*Proof.*  $K - M_t$  is a local martingale with no negative jumps. So it follows from Lemma 2.1 that  $(K - M_t)^+$  is a local submartingale of class  $(\Sigma)$ . Since  $(M_t^-)$  is of class (D),  $(K - M_t)^+$  is of class  $(\Sigma D)$  and (3.5) follows from Theorem 3.1 by noting that  $g_K = \sup\{t : (K - M_t)^+ = 0\}$ .  $\square$

**Remark 3.5** If  $K$  is a constant and  $(M_t)$  a local martingale with no negative jumps such that  $(M_t^+)$  is of class (D), one can apply Corollary 3.4 to  $-K$ ,  $(-M_t)$  and  $g_K = \sup\{t : M_t \leq K\}$ . This gives

$$(M_T - K)^+ = \mathbb{E}[(M_\infty - K)^+ 1_{\{g_K \leq T\}} \mid \mathcal{F}_T] \quad (3.6)$$

for all stopping times  $T$ . In particular, if  $M_\infty = m \in \mathbb{R}$ , then

$$(M_T - K)^+ = (m - K)^+ \mathbb{P}[g_K \leq T \mid \mathcal{F}_T]. \quad (3.7)$$

However, if for instance,  $M_t = \exp(B_t - t/2)$  for a Brownian motion  $(B_t)$ , the assumptions of Corollary 3.4 are satisfied but  $(M_t^+)$  is not of class (D). So even though  $M_\infty = 0$ , formula (3.7) does not hold. Indeed, for  $K \geq 0$  and  $T = t \in \mathbb{R}_+ \setminus \{0\}$ , the right-hand side is zero but  $\mathbb{P}[(M_t - K)^+ > 0] > 0$ . For a more detailed discussion of this case, we refer to Section 6 in Madan et al. [13].

The following extension of Corollary 3.4 has been proved by Profeta et al. [23] with methods from the theory of enlargement of filtrations. We can deduce it under slightly weaker assumptions from Lemma 2.2 and Theorem 3.1.

**Corollary 3.6** (Profeta–Roynette–Yor [23])

Let  $K^1, \dots, K^n$  be constants and  $(M_t^1), \dots, (M_t^n)$  local martingales that are bounded from below and have no positive jumps. Assume  $[M^i, M^j]_t = 0$  for  $i \neq j$  and denote  $g^i = \sup \{t : M_t^i \geq K^i\}$ . Then

$$\prod_{i=1}^n (K^i - M_T^i)^+ = \mathbb{E} \left[ \prod_{i=1}^n (K^i - M_\infty^i)^+ 1_{\{g^i \leq T\}} \mid \mathcal{F}_T \right], \quad (3.8)$$

for every stopping time  $T$ . In particular, if  $M_\infty^i = m^i \in \mathbb{R}$  for all  $i = 1, \dots, n$ , then

$$\prod_{i=1}^n (K^i - M_T^i)^+ = \prod_{i=1}^n (K^i - m^i)^+ \mathbb{P} \left[ \bigvee_{i=1}^n g^i \leq T \mid \mathcal{F}_T \right]$$

*Proof.* By Lemma 2.1,  $X_t^i = (K^i - M_t^i)^+$  are local submartingales of class  $(\Sigma)$  such that  $[X^i, X^j]_t = 0$  for  $i \neq j$ . So we obtain from Lemma 2.2 that  $\prod_{i=1}^n X_t^i$  is again of class  $(\Sigma)$ , which since all  $(M_t^i)$  are bounded from below, is bounded. Now (3.8) follows from Theorem 3.1.  $\square$

**Remark 3.7** If  $(X_t)$  satisfies the assumptions of part (2) of Theorem 3.1 or Corollary 3.2, then there exists a random variable  $X_\infty$  such that  $X_t \rightarrow X_\infty$  almost everywhere on the set  $\{L < \infty\}$ , and one has

$$X_t = \mathbb{E} [X_\infty 1_{\{L \leq t\}} \mid \mathcal{F}_t] \quad \text{for all } t \geq 0. \quad (3.9)$$

In particular, the whole process  $(X_t)$  can be recovered from  $X_\infty$  and  $L$ . If  $(X_t)$  is non-negative, equation (3.9) can be rewritten as

$$\mathbb{E} [1_{F_t} X_t] = \mathcal{Q}[F_t \cap \{g \leq t\}] \quad \text{for every } t \geq 0 \text{ and all } F_t \in \mathcal{F}_t, \quad (3.10)$$

where  $g = L$  and  $\mathcal{Q}$  is the  $\sigma$ -finite measure given by  $d\mathcal{Q}/d\mathbb{P} = X_\infty$ . Madan et al. [13] raised the question for which non-negative submartingales  $(X_t)$  is it possible to find a random time  $g$  and a  $\sigma$ -finite measure  $\mathcal{Q}$  such that (3.10) holds. It turns out that if  $(X_t)$  satisfies (3.10), where  $g$  is the end of an optional set with the property

$$\mathbb{P}[g = T] = 0 \quad \text{for all stopping times } T \quad (3.11)$$

and  $\mathcal{Q}$  is of the form  $d\mathcal{Q}/d\mathbb{P} = X$  for an integrable random variable  $X > 0$ , then  $(X_t)$  is a submartingale of class  $(\Sigma D)$  with  $X_\infty = \mathbb{E}[X \mid \bigvee_t \mathcal{F}_t]$  and  $L = g$ . For  $X = 1$ , this holds because in this case,  $(X_t)$  can be written as  $X_t = 1 - Z_t^g$  for the Azéma supermartingale  $Z_t^g = \mathbb{P}[g > t \mid \mathcal{F}_t]$ . Moreover, it follows from (3.11) that  $g$  is the end of a predictable set and the finite variation part  $A_t^g$  of  $Z_t^g$  is continuous. It is shown in Azéma [1] that under these circumstances, one has  $\int_0^t 1_{\{Z_u^g \neq 1\}} dA_u^g = 0$  and  $g = \sup \{t : Z_t^R = 1\}$ , which means that  $(X_t)$  is a submartingale of class  $(\Sigma D)$  with  $L = g$ . The case of general integrable  $X > 0$  follows from an application of the optional section theorem (see page 136 of Dellacherie et al. [8]). A more thorough discussion of this problem, also treating the case where  $\mathcal{Q}$  is a  $\sigma$ -finite measure, is the subject of the paper Najnudel and Nikeghbali [16].

### 3.2 Stochastic integral representations and conditional distributions

We here use Lemma 2.3 to derive representation results for non-negative processes  $(X_t)$  of class  $(\Sigma)$  and Borel functions  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  such that  $f(A_t)X_t$  converges to 1 as  $t \rightarrow \infty$ . In Section 4 they will be applied in situations where  $f(A_t)X_t$  can be stopped with a stopping time  $R$  such that  $f(A_R)X_R = 1$ .

**Theorem 3.8** *Let  $(X_t)$  be a non-negative process of class  $(\Sigma)$  and  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  a locally bounded Borel function such that the process  $f(A_t)X_t$  is of class  $(D)$  and  $f(A_t)X_t \rightarrow 1$  almost surely. Denote  $F(x) = \int_0^x f(y)dy$ .*

- (1) *If  $F(\infty) < \infty$ , then  $A_t = 0$  and  $f(0)X_t = 1$  for all  $t \geq 0$ .*
- (2) *If  $F(\infty) = \infty$ , then  $L < \infty$ ,  $A_L = A_\infty < \infty$  and  $X_t \rightarrow X_\infty$  almost surely for a random variable  $X_\infty > 0$ . Moreover, for every stopping time  $T$  one has*

$$f(A_T)X_T = \mathbb{P}[L \leq T \mid \mathcal{F}_T] \quad (3.12)$$

and for all Borel functions  $h : \mathbb{R}_+ \rightarrow \mathbb{R}$  satisfying

$$\int_0^\infty |h(y)|e^{-F(y)}dF(y) < \infty, \quad (3.13)$$

$$\mathbb{E}[h(A_\infty) \mid \mathcal{F}_T] = h(0)f(0)X_0 + h^F(0)(1 - f(0)X_0) + \int_0^T (h - h^F)(A_u)f(A_u)dN_u \quad (3.14)$$

$$= h(A_T)f(A_T)X_T + h^F(A_T)(1 - f(A_T)X_T), \quad (3.15)$$

where

$$h^F(x) = e^{F(x)} \int_x^\infty h(y)e^{-F(y)}dF(y), \quad x \geq 0.$$

In particular, conditioned on  $\mathcal{F}_T$ , the law of  $A_\infty$  is given by

$$\mathbb{P}[A_\infty > x \mid \mathcal{F}_T] = 1_{\{A_T > x\}} + 1_{\{A_T \leq x\}}(1 - f(A_T)X_T)e^{F(A_T) - F(x)}, \quad x \geq 0. \quad (3.16)$$

*Proof.* Since  $f(A_t)X_t \rightarrow 1$  almost surely, one has  $L < \infty$  and  $A_t = A_{t \wedge L}$ . In particular,  $A_L = A_\infty$  and there exists a random variable  $X_\infty > 0$  such that  $X_t \rightarrow X_\infty$  almost surely. By Lemma 2.3, the process  $f(A_t)X_t$  is of class  $(\Sigma D)$ . Hence it follows from Theorem 3.1 that for every stopping time  $T$ ,

$$f(A_T)X_T = \mathbb{P}[L^f \leq T \mid \mathcal{F}_T], \quad (3.17)$$

where  $L^f = \sup\{t : f(A_t)X_t = 0\}$ . However, since  $f(A_t)X_t \rightarrow 1$ , one has  $L = L^f$  and (3.12) follows from (3.17). Now let  $h : \mathbb{R}_+ \rightarrow \mathbb{R}$  be a bounded Borel function. Then the function

$$h^F(x) = e^{F(x)} \int_x^\infty h(y)e^{-F(y)}dF(y)$$

is bounded as well. Note that  $\Phi(x) = h^F(0) - h^F(x)$  is of the form  $\Phi(x) = \int_0^x \varphi(y)dy$  for  $\varphi = (h - h^F)f$ . So one obtains from Lemma 2.3 that  $\varphi(A_t)X_t$  is of class  $(\Sigma D)$  and

$$\varphi(0)X_0 + \int_0^T \varphi(A_u)dN_u = \varphi(A_T)X_T - \Phi(A_T) = \mathbb{E}[\varphi(A_\infty)X_\infty - \Phi(A_\infty) \mid \mathcal{F}_T]$$

for every stopping time  $T$ . This yields

$$\mathbb{E}[h(A_\infty) \mid \mathcal{F}_T] = h(0)f(0)X_0 + h^F(0)(1 - f(0)X_0) + \int_0^T (h - h^F)(A_u)f(A_u)dN_u \quad (3.18)$$

$$= h(A_T)f(A_T)X_T + h^F(A_T)(1 - f(A_T)X_T). \quad (3.19)$$

(3.19) applied to  $h \equiv 1$  gives  $e^{F(A_t)-F(\infty)}(1 - f(A_t)X_t) = 0$  for all  $t \geq 0$ . Hence, for  $F(\infty) < \infty$ , one must have  $f(A_t)X_t = 1$  for all  $t \geq 0$ , and (1) follows. If  $F(\infty) = \infty$ , then formula (3.19) is equivalent to (3.16), which shows that condition (3.13) implies  $\mathbb{E}[|h(A_\infty)| \mid \mathcal{F}_T] < \infty$ . So both equations (3.18) and (3.19) extend from bounded  $h$  to functions that satisfy (3.13).  $\square$

Theorem 3.8 will allow us to obtain general results on drawdown and relative drawdown processes of local martingales in Section 4. But to extend these results to diffusions we will need the following generalization of Theorem 3.8. The proof is similar but involves an additional approximation argument.

**Theorem 3.9** *Let  $(X_t)$  be a non-negative process of class  $(\Sigma)$  and  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  a Borel function for which there exists an increasing sequence  $(a_n)_{n \in \mathbb{N}}$  in  $(0, \infty)$  such that  $f1_{[0, a_n]}$  is bounded for all  $n \in \mathbb{N}$ , and  $f(x) = 0$  for  $x \geq a = \lim_{n \rightarrow \infty} a_n$ . Denote  $F(x) = \int_0^x f(y)dy$  and assume that the process  $f(A_t)X_t$  is of class  $(D)$  and  $f(A_t)X_t \rightarrow 1$  almost surely.*

- (1) *If  $F(a) < \infty$ , then  $A_t = 0$  and  $f(0)X_t = 1$  for all  $t \geq 0$ .*
- (2) *If  $F(a) = \infty$ , then  $L < \infty$ ,  $A_L = A_\infty < a$  and  $X_t \rightarrow X_\infty$  almost surely for a random variable  $X_\infty > 0$ . Moreover, for every stopping time  $T$  one has*

$$f(A_T)X_T = \mathbb{P}[L \leq T \mid \mathcal{F}_T] \quad (3.20)$$

and for all Borel functions  $h : [0, a) \rightarrow \mathbb{R}$  satisfying

$$\int_0^a |h(y)|e^{-F(y)}dF(y) < \infty, \quad (3.21)$$

$$\mathbb{E}[h(A_\infty) \mid \mathcal{F}_T] = h(0)f(0)X_0 + h^F(0)(1 - f(0)X_0) + \int_0^T (h - h^F)(A_u)f(A_u)dN_u \quad (3.22)$$

$$= h(A_T)f(A_T)X_T + h^F(A_T)(1 - f(A_T)X_T), \quad (3.23)$$

where

$$h^F(x) = e^{F(x)} \int_x^a h(y)e^{-F(y)}dF(y), \quad 0 \leq x < a.$$

In particular, conditioned on  $\mathcal{F}_T$ , the law of  $A_\infty$  is given by

$$\mathbb{P}[A_\infty > x \mid \mathcal{F}_T] = 1_{\{A_T > x\}} + 1_{\{A_T \leq x\}}(1 - f(A_T)X_T)e^{F(A_T)-F(x)}, \quad x \geq 0. \quad (3.24)$$

*Proof.* It follows as in the proof of Theorem 3.8 that  $L < \infty$ ,  $A_L = A_\infty$  and  $X_t \rightarrow X_\infty$  almost surely for a random variable  $X_\infty > 0$ . Now set  $f^n = f \wedge n$  and  $F^n(x) = \int_0^x f^n(y)dy$ ,  $n \in \mathbb{N}$ . It follows from Lemma 2.3 that the processes  $f^n(A_t)X_t$ ,  $n \in \mathbb{N}$ , are of class  $(\Sigma D)$ . So one obtains from Theorem 3.1 that

$$f^n(A_T)X_T = \mathbb{E} [f^n(A_\infty)X_\infty 1_{\{L^n \leq T\}} | \mathcal{F}_T] \quad \text{for every stopping time } T, \quad (3.25)$$

where  $L^n = \sup \{t : f^n(A_t)X_t = 0\}$ . The fact that  $f^n(A_t)X_t \rightarrow f^n(A_\infty)X_\infty > 0$  entails that  $L = L^n$ , and one obtains (3.20) from (3.25) by letting  $n$  tend to  $\infty$ . Now let  $h : \mathbb{R}_+ \rightarrow \mathbb{R}$  be a bounded Borel function. Then the functions

$$h^F(x) = e^{F(x)} \int_x^\infty h(y)e^{-F(y)}dF(y) \quad \text{and} \quad h^n(x) = e^{F^n(x)} \int_x^\infty h(y)e^{-F^n(y)}dF^n(y), \quad n \in \mathbb{N},$$

are bounded too.  $\Phi^n(x) = h^n(0) - h^n(x)$  can be written as  $\Phi^n(x) = \int_0^t \varphi^n(y)dy$  for  $\varphi^n = (h - h^n)f^n$ . So it follows from Lemma 2.3 that  $\varphi^n(A_t)X_t$  is of class  $(\Sigma D)$  and

$$\varphi^n(0)X_0 + \int_0^T \varphi^n(A_u)dN_u = \varphi^n(A_T)X_T - \Phi^n(A_T) = \mathbb{E} [\varphi^n(A_\infty)X_\infty - \Phi^n(A_\infty) | \mathcal{F}_T]$$

for every stopping time  $T$ . In the limit  $n \rightarrow \infty$ , this gives

$$\mathbb{E} [h(A_\infty) | \mathcal{F}_T] = h(0)f(0)X_0 + h^F(0)(1 - f(0)X_0) + \int_0^T (h - h^F)(A_u)f(A_u)dN_u \quad (3.26)$$

$$= h(A_T)f(A_T)X_T + h^F(A_T)(1 - f(A_T)X_T). \quad (3.27)$$

For  $h = 1_{[a, \infty)}$  and  $T = 0$ , the equality between the first and third term reduces to  $\mathbb{P}[A_\infty \geq a] = 0$ . So  $A_\infty < a$  and (3.27) applied to  $h \equiv 1$  gives  $e^{F(A_t)-F(a)}(1 - f(A_t)X_t) = 0$  for all  $t \geq 0$ . Hence, for  $F(a) < \infty$ , one must have  $f(A_t)X_t = 1$  for all  $t \geq 0$ , and (1) follows. If  $F(a) = \infty$ , formula (3.27) is equivalent to (3.24). This shows that condition (3.21) implies  $\mathbb{E} [|h(A_\infty)| | \mathcal{F}_T] < \infty$ , and both equations (3.26) and (3.27) extend from bounded  $h$  to functions satisfying (3.21).  $\square$

## 4 Drawdown and relative drawdown

### 4.1 The local martingale case

We first consider a local martingale  $(M_t)$  starting at  $m \in \mathbb{R}$  such that the running supremum  $(\overline{M}_t)$  is continuous. Then the drawdown process  $DD_t = \overline{M}_t - M_t$  is a non-negative local submartingale of class  $(\Sigma)$  with decomposition  $(m - M_t) + (\overline{M}_t - m)$ . Moreover, if  $m > 0$ , then the relative drawdown process  $rDD_t = DD_t/\overline{M}_t$  is well-defined, and by Lemma 2.3, also a non-negative local submartingale of class  $(\Sigma)$ . As a consequence of the results of Subsection 3.1 one obtains the following

**Proposition 4.1** *Assume that  $(M_t^-)$  is of class (D) and denote  $g = \sup \{t : M_t = \overline{M}_t\}$ . Then there exists a random variable  $DD_\infty$  such that  $DD_t \rightarrow DD_\infty$  almost everywhere on  $\{g < \infty\}$  and*

$$DD_T = \mathbb{E} [DD_\infty 1_{\{g \leq T\}} | \mathcal{F}_T] \quad (4.1)$$

for every stopping time  $T < \infty$ . Moreover, if  $m > 0$ , then there exists an integrable random variable  $rDD_\infty$  such that  $rDD_t \rightarrow rDD_\infty$  almost surely as well as in  $L^1$  and

$$rDD_T = \mathbb{E} [rDD_\infty 1_{\{g \leq T\}} | \mathcal{F}_T] \quad (4.2)$$

$$= - \int_0^T \frac{dM_u}{\overline{M}_u} + \log(\overline{M}_T) - \log(m) \quad (4.3)$$

$$= \mathbb{E} [rDD_\infty - \log(\overline{M}_\infty) | \mathcal{F}_T] + \log(\overline{M}_T). \quad (4.4)$$

for all stopping times  $T$ .

*Proof.*  $DD_t$  is a process of class  $(\Sigma)$  with  $L = g$  that satisfies the assumptions of Corollary 3.2. This shows (4.1). Moreover, if  $m > 0$ , then  $rDD_t$  is a process of class  $(\Sigma D)$ , and (4.2)–(4.4) follow from Theorem 3.1 and Lemma 2.3.  $\square$

In the following we study the situation where  $(M_t)$  is stopped with a stopping time of the form  $T_\lambda = \inf \{t : M_t \leq \lambda(\overline{M}_t)\}$  for a Borel function  $\lambda : [m, \infty) \rightarrow \mathbb{R}$  satisfying  $\lambda(x) < x$  for all  $x \geq m$ . Denote

$$g_\lambda = \sup \{t \leq T_\lambda : M_t = \overline{M}_t\}, \quad \Lambda(x) = \int_m^x \frac{dy}{y - \lambda(y)} \quad (4.5)$$

and notice that  $\Lambda$  is a well-defined increasing function from  $[m, \infty)$  to  $[0, \infty]$ .

**Proposition 4.2** *Assume the function  $1/(x - \lambda(x))$  is locally bounded on  $[m, \infty)$  and*

$$\overline{M}_{T_\lambda} < \infty \quad \text{and} \quad M_{T_\lambda} = \lambda(\overline{M}_{T_\lambda}). \quad (4.6)$$

Then  $\Lambda(\infty) = \infty$ ,  $g_\lambda < T_\lambda$ ,  $M_{T_\lambda} < \overline{M}_{g_\lambda} = \overline{M}_{T_\lambda}$  and

$$\mathbb{P}[g_\lambda \leq T | \mathcal{F}_T] = \frac{\overline{M}_T - M_T}{\overline{M}_T - \lambda(\overline{M}_T)} \quad (4.7)$$

for every stopping time  $T \leq T_\lambda$ . Moreover, for all Borel functions  $h : [m, \infty) \rightarrow \mathbb{R}$  satisfying

$$\int_m^\infty |h(y)| e^{-\Lambda(y)} d\Lambda(y) < \infty$$

one has

$$\mathbb{E} [h(\overline{M}_{T_\lambda}) | \mathcal{F}_T] = h^\Lambda(m) + \int_0^T \frac{h^\Lambda(\overline{M}_u) - h(\overline{M}_u)}{\overline{M}_u - \lambda(\overline{M}_u)} dM_u \quad (4.8)$$

$$= \frac{h(\overline{M}_T)(\overline{M}_T - M_T) + h^\Lambda(\overline{M}_T)(M_T - \lambda(\overline{M}_T))}{\overline{M}_T - \lambda(\overline{M}_T)}, \quad (4.9)$$

where

$$h^\Lambda(x) = e^{\Lambda(x)} \int_x^\infty h(y) e^{-\Lambda(y)} d\Lambda(y), \quad x \geq m. \quad (4.10)$$

In particular,

$$\mathbb{P}[\overline{M}_{T_\lambda} > x | \mathcal{F}_T] = 1_{\{\overline{M}_T > x\}} + 1_{\{\overline{M}_T \leq x\}} \frac{M_T - \lambda(\overline{M}_T)}{\overline{M}_T - \lambda(\overline{M}_T)} e^{\Lambda(\overline{M}_T) - \Lambda(x)} \quad \text{for } x \geq m. \quad (4.11)$$

*Proof.*  $X_t = \overline{M}_{t \wedge T_\lambda} - M_{t \wedge T_\lambda}$  is a non-negative process of class  $(\Sigma)$  starting at 0 with decomposition  $(m - M_{t \wedge T_\lambda}) + (\overline{M}_{t \wedge T_\lambda} - m)$ , and

$$f(x) = \frac{1}{x + m - \lambda(x + m)}$$

is a non-negative locally bounded Borel function on  $\mathbb{R}_+$  with  $F(x) = \int_0^x f(y)dy = \Lambda(x + m)$ . By Lemma 2.3, the process

$$f(\overline{M}_{t \wedge T_\lambda} - m)X_t = \frac{\overline{M}_{t \wedge T_\lambda} - M_{t \wedge T_\lambda}}{\overline{M}_{t \wedge T_\lambda} - \lambda(\overline{M}_{t \wedge T_\lambda})}$$

is of class  $(\Sigma)$ . Moreover, it follows from condition (4.6) that it takes values in  $[0, 1]$  and converges to 1 almost surely. Since  $X_0 = 0$ , one obtains from Theorem 3.8 that  $\Lambda(\infty) = \int_0^\infty f(x)dx = \infty$ ,  $g_\lambda < T_\lambda$ ,  $M_{T_\lambda} < \overline{M}_{g_\lambda} = \overline{M}_{T_\lambda}$  and

$$\mathbb{P}[g_\lambda \leq T \mid \mathcal{F}_T] = \frac{\overline{M}_T - M_T}{\overline{M}_T - \lambda(\overline{M}_T)}$$

for all stopping times  $T \leq T_\lambda$ . Formulas (4.8)–(4.11) follow from Theorem 3.8 applied to the function

$$\tilde{h}(x) = h(x + m), \quad x \geq 0.$$

□

**Remark 4.3** If  $(M_t)$  is a non-negative local martingale starting at 1 such that  $(\overline{M}_t)$  is continuous and  $M_t \rightarrow 0$  almost surely, then formula (4.11) with  $T = 0$  and  $\lambda \equiv 0$  yields that  $1/\overline{M}_\infty$  is uniformly distributed on the interval  $(0, 1)$ . This is Doob's maximal identity, which has been studied in depth by Mansuy and Yor [15] and Nikeghbali and Yor [20].

The next proposition gives sufficient conditions for assumption (4.6) to hold.

**Proposition 4.4** *Assume  $(M_t)$  is continuous. Then both of the following two conditions imply condition (4.6).*

- (1)  $m > 0$ ,  $M_t \rightarrow 0$  almost surely and  $\lambda(x) \geq 0$  for all  $x \geq m$
- (2)  $\overline{M}_\infty = \Lambda(\infty) = \infty$

*Proof.* Under condition (1) one has  $\overline{M}_{T_\lambda} \leq \overline{M}_\infty < \infty$  and

$$\frac{\overline{M}_t - M_t}{\overline{M}_t - \lambda(\overline{M}_t)} \geq \frac{\overline{M}_t - M_t}{\overline{M}_t} \rightarrow 1 \text{ almost surely,}$$

which shows that  $M_{T_\lambda} = \lambda(\overline{M}_{T_\lambda})$ .

If condition (2) holds, then  $\int_0^\infty f(x)dx = \infty$  for the function

$$f(x) = \frac{1}{x + m - \lambda(x + m)}.$$

Since  $X_t = \overline{M}_t - M_t$  is a non-negative continuous process of class  $(\Sigma)$  with decomposition  $(m - M_t) + (\overline{M}_t - m)$ , it follows from Corollary 2.4 that  $\mathbb{P}[f(\overline{M}_t - m)X_t < 1 \text{ for all } t] = 0$ . This implies  $T_\lambda < \infty$ ,  $\overline{M}_{T_\lambda} < \infty$  and  $M_{T_\lambda} = \lambda(\overline{M}_{T_\lambda})$ . □



Several authors have studied the distribution of the maximum drawdown  $\sup_{0 \leq t \leq T} DD_t$  in the case where  $T$  is a constant and  $(M_t)$  a Brownian motion (with or without drift); see, for instance, Berger and Whitt [5], Douady et al. [9], Graversen and Shiryaev [11], Magdon-Ismail et al. [14]. With the methods developed here we can derive conditional distributions of maximum drawdowns  $\sup_{T \leq t \leq R} DD_t$  and maximum relative drawdowns  $\sup_{T \leq t \leq R} rDD_t$  when  $T \leq R$  are suitable stopping times and  $(M_t)$  is a continuous local martingale. In the next subsection we will generalize these results to the diffusion case.

**Proposition 4.5** *Assume  $(M_t)$  is continuous with  $m > 0$  and  $M_t \rightarrow 0$  almost surely. Let  $T < \infty$  be a stopping time and  $K$  an  $\mathcal{F}_T$ -measurable random variable such that  $0 \leq K < M_T$ . Denote  $T_K = \inf \{t \geq T : M_t = K\}$ . Then one has for all  $x \geq 0$ ,*

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t > x \mid \mathcal{F}_T \right] = 1_{\{\overline{M}_T - K > x\}} + 1_{\{\overline{M}_T - K \leq x\}} \frac{M_T - K}{x}. \quad (4.12)$$

If in addition,  $m > 0$ , then

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t > x \mid \mathcal{F}_T \right] = 1_{\{1 - K/\overline{M}_T > x\}} + 1_{\{1 - K/\overline{M}_T \leq x < 1\}} \left( \frac{M_T - K}{K} \right) \left( \frac{1 - x}{x} \right). \quad (4.13)$$

*Proof.* Let us first assume  $T = 0$  and  $K$  is a constant. Then, by Proposition 4.4,  $(M_t)$  with  $\lambda \equiv K$  fulfills the assumptions of Proposition 4.2. Moreover,

$$\sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t = \overline{M}_{T_K} - K \quad \text{and} \quad \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t = 1 - K/\overline{M}_{T_K}.$$

So (4.12) and (4.13) can be deduced from formula (4.11). In the general case, the proposition follows by considering the process  $\tilde{M}_t = M_{T+t}$  in the filtration  $\tilde{\mathcal{F}}_t = \mathcal{F}_{T+t}$  and conditioning on  $\mathcal{F}_T$ .  $\square$

**Proposition 4.6** *Assume  $(M_t)$  is continuous and  $\overline{M}_\infty = \infty$ . Let  $T < \infty$  be a stopping time and  $K$  a  $[m, \infty]$ -valued  $\mathcal{F}_T$ -measurable random variable such that  $\overline{M}_T < K \leq \overline{M}_\infty$ . Denote*

$$T_K = \inf \{t \geq T : M_t = K\}$$

. Then one has for all  $x \geq 0$ ,

$$\mathbb{P}[M_t \geq \lambda(\overline{M}_t) \text{ for all } t \in [T, T_K] \cap \mathbb{R}_+ \mid \mathcal{F}_T] = \left( \frac{M_T - \lambda(\overline{M}_T)}{\overline{M}_T - \lambda(\overline{M}_T)} \right)^+ e^{\Lambda(\overline{M}_T) - \Lambda(K)}. \quad (4.14)$$

*Proof.*  $X_t = \overline{M}_t - M_t$  is a non-negative process of class  $(\Sigma)$  with decomposition  $(m - M_t) + (\overline{M}_t - m)$  and

$$f(x) = \frac{1}{x + m - \lambda(x + m)}$$

a Borel function from  $\mathbb{R}_+$  to  $\mathbb{R}_+$ . Note that  $M_t \geq \lambda(\overline{M}_t)$  is equivalent to  $f(\overline{M}_t - m)X_t \leq 1$  and  $T_k = \inf \{\overline{M}_t - m = K - m\}$ . Therefore, it follows from Corollary 2.4 that

$$\begin{aligned} \mathbb{P}[M_t \geq \lambda(\overline{M}_t) \text{ for all } t \in [T, T_K] \cap \mathbb{R}_+ \mid \mathcal{F}_T] &= \left( \frac{M_T - \lambda(\overline{M}_T)}{\overline{M}_T - \lambda(\overline{M}_T)} \right)^+ \exp \left( - \int_{\overline{M}_T - m}^{K - m} f(x) dx \right) \\ &= \left( \frac{M_T - \lambda(\overline{M}_T)}{\overline{M}_T - \lambda(\overline{M}_T)} \right)^+ \exp (\Lambda(\overline{M}_T) - \Lambda(K)). \end{aligned}$$

□

**Corollary 4.7** *Under the assumptions of Proposition 4.6 one has for all  $x \geq 0$ ,*

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t \leq x \mid \mathcal{F}_T \right] = 1_{\{x > 0\}} \left( 1 - \frac{DD_T}{x} \right)^+ \exp \left( \frac{\overline{M}_T - K}{x} \right), \quad (4.15)$$

and, provided that  $m > 0$ ,

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t \leq x \mid \mathcal{F}_T \right] = 1_{\{x > 0\}} \left( 1 - \frac{rDD_T}{x} \right)^+ \left( \frac{\overline{M}_T}{K} \right)^{1/x}. \quad (4.16)$$

*Proof.* First assume  $x > 0$ . Then formula (4.15) follows from Proposition 4.6 applied to the function  $\lambda(y) = y - x$  and (4.16) is obtained by applying Proposition 4.6 with  $\lambda(y) = (1 - x)y$ . Since  $\overline{M}_T < K$ , one has

$$\exp \left( \frac{\overline{M}_T - K}{x} \right) \rightarrow 0 \quad \text{and} \quad \left( \frac{\overline{M}_T}{K} \right)^{1/x} \rightarrow 0 \quad \text{almost surely for } x \downarrow 0.$$

This shows that

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t = 0 \mid \mathcal{F}_T \right] = \mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t = 0 \mid \mathcal{F}_T \right] = 0.$$

□

## 4.2 The diffusion case

The results of Subsection 4.1 can be extended to stochastic processes that can be turned into local martingales through a strictly increasing continuous transformation. To do that we here consider a stochastic process  $(Y_t)$  taking values in an interval  $I \subset \mathbb{R}$  which starts at a constant  $y_0 \in I$  such that the supremum process  $(\overline{Y}_t)$  is continuous and there exists a strictly increasing continuous function  $s : I \rightarrow \mathbb{R}$  making  $s(Y_t)$  a local martingale. Our main example is a diffusion of the form

$$dY_t = \mu(Y_t)dt + \sigma(Y_t)dB_t, \quad Y_0 = y_0 \in I, \quad (4.17)$$

where  $(B_t)$  is a Brownian motion and  $\mu, \sigma : I \rightarrow \mathbb{R}$  are deterministic functions such that

$$\gamma(x) = 2 \int_{y_0}^x \frac{\mu(y)}{\sigma^2(y)} dy \quad \text{and} \quad \int_{y_0}^x e^{-\gamma(y)} dy \quad \text{are finite for all } x \in I.$$

Then  $s$  can be chosen as  $s(x) = c + d \int_{y_0}^x e^{-\gamma(y)} dy$  for arbitrary constants  $c \in \mathbb{R}$  and  $d > 0$ . For instance, if  $Y_t = B_t + bt$  for  $b \in \mathbb{R} \setminus \{0\}$ , then  $I = \mathbb{R}$  and  $s$  can be chosen as  $s(x) = -\text{sign}(b)e^{-2bx}$ . Or if  $(Y_t)$  is a Bessel process of dimension  $\delta = 2(1 - \nu) > 2$  starting at  $y_0 > 0$ , then one can choose  $I = (0, \infty)$  and  $s(x) = -x^{2\nu}$ .

Denote the drawdown process  $\overline{Y}_t - Y_t$  by  $DD_t$  and if  $y_0 > 0$ , the relative drawdown  $DD_t/\overline{Y}_t$  by  $rDD_t$ . As in Subsection 4.1 we consider a stopping time of the form

$$T_\lambda = \inf \{t : Y_t \leq \lambda(\overline{Y}_t)\}$$

for a Borel function  $\lambda$ . But this time we assume that  $\lambda$  maps  $[y_0, \infty) \cap I$  to the closure  $\bar{I}$  of  $I$  in  $[-\infty, \infty]$  such that  $\lambda(x) < x$  for all  $x \in [y_0, \sup I)$ . Extend  $s$  continuously to  $s : \bar{I} \rightarrow [-\infty, \infty]$  and denote

$$g_\lambda = \sup \{t \leq T_\lambda : Y_t = \bar{Y}_t\} \quad \text{and} \quad \Lambda(x) = \int_{y_0}^x \frac{ds(y)}{s(y) - s \circ \lambda(y)}. \quad (4.18)$$

Then  $\Lambda$  is a well-defined increasing function from  $[y_0, \sup I)$  to  $[0, \infty]$ . The following result generalizes Proposition 4.2:

**Proposition 4.8** *Let  $(a_n)_{n \in \mathbb{N}}$  be an increasing sequence in  $(y_0, \sup I)$  and  $(\varepsilon_n)_{n \in \mathbb{N}}$  a decreasing sequence in  $(0, \infty)$  such that  $\lambda(x) \leq x - \varepsilon_n$  for  $y_0 \leq x \leq a_n$ . Denote  $a = \lim_{n \rightarrow \infty} a_n \in (y_0, \sup I]$  and assume that*

$$\bar{Y}_{T_\lambda} < a \quad \text{and} \quad Y_{T_\lambda} = \lambda(\bar{Y}_{T_\lambda}). \quad (4.19)$$

Then  $\Lambda(a) = \infty$ ,  $g_\lambda < T_\lambda$ ,  $Y_{T_\lambda} < \bar{Y}_{g_\lambda} = \bar{Y}_{T_\lambda}$  and

$$\mathbb{P}[g_\lambda \leq T \mid \mathcal{F}_T] = \frac{s(\bar{Y}_T) - s(Y_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)} \quad (4.20)$$

for every stopping time  $T \leq T_\lambda$ . Moreover, for all Borel functions  $h : [y_0, a) \rightarrow \mathbb{R}$  satisfying

$$\int_{y_0}^a |h(y)| e^{-\Lambda(y)} d\Lambda(y) < \infty$$

one has

$$\mathbb{E} [h(\bar{Y}_{T_\lambda}) \mid \mathcal{F}_T] = h^\Lambda(y_0) + \int_0^T \frac{h^\Lambda(\bar{Y}_u) - h^\Lambda(\bar{Y}_u)}{s(\bar{Y}_u) - s \circ \lambda(\bar{Y}_u)} ds(Y_u) \quad (4.21)$$

$$= \frac{h(\bar{Y}_T)[s(\bar{Y}_T) - s(Y_T)] + h^\Lambda(\bar{Y}_T)[s(Y_T) - s \circ \lambda(\bar{Y}_T)]}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)}, \quad (4.22)$$

where

$$h^\Lambda(x) = e^{\Lambda(x)} \int_x^a h(y) e^{-\Lambda(y)} d\Lambda(y), \quad x \geq y_0.$$

In particular,

$$\mathbb{P}[\bar{Y}_{T_\lambda} > x \mid \mathcal{F}_T] = 1_{\{\bar{Y}_T > x\}} + 1_{\{\bar{Y}_T \leq x\}} \frac{s(Y_T) - s \circ \lambda(\bar{Y}_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)} e^{\Lambda(\bar{Y}_T) - \Lambda(x)} \quad \text{for } x \geq y_0. \quad (4.23)$$

*Proof.*  $X_t = s(\bar{Y}_{t \wedge T_\lambda}) - s(Y_{t \wedge T_\lambda})$  is a non-negative process of class  $(\Sigma)$  starting at 0 with decomposition  $(s(y_0) - s(Y_{t \wedge T_\lambda})) + (s(\bar{Y}_{t \wedge T_\lambda}) - s(y_0))$ . The function  $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  given by

$$f(x) = 1_{\{x < s(a) - s(y_0)\}} \frac{1}{x + s(y_0) - s \circ \lambda \circ s^{-1}(x + s(y_0))}$$

satisfies the assumptions of Theorem 3.9 with  $\tilde{a}_n = s(a_n) - s(y_0)$  and  $\tilde{a} = s(a) - s(y_0)$  instead of  $a_n$  and  $a$ . Assumption (4.19) guarantees that the process

$$f(s(\bar{Y}_{t \wedge T_\lambda}) - s(y_0)) X_t = 1_{\{\bar{Y}_{t \wedge T_\lambda} < a\}} \frac{s(\bar{Y}_{t \wedge T_\lambda}) - s(Y_{t \wedge T_\lambda})}{s(\bar{Y}_{t \wedge T_\lambda}) - s \circ \lambda(\bar{Y}_{t \wedge T_\lambda})}$$

takes values in  $[0, 1]$  and converges to 1 almost surely. So it follows from Theorem 3.9 that  $\Lambda(a) = \int_0^{\tilde{a}} f(x)dx = \infty$ ,  $g_\lambda < T_\lambda$ ,  $Y_{T_\lambda} < \bar{Y}_{g_\lambda} = \bar{Y}_{T_\lambda}$  and

$$\mathbb{P}[g_\lambda \leq T \mid \mathcal{F}_T] = \frac{s(\bar{Y}_T) - s(Y_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)}$$

for all stopping times  $T \leq T_\lambda$ . Formulas (4.21)–(4.23) follow from Theorem 3.9 applied to the function

$$\tilde{h}(x) = h(s^{-1}(x + s(y_0))), \quad 0 \leq x < \tilde{a}.$$

□

The generalization of Proposition 4.4 to the context of the current subsection looks as follows:

**Proposition 4.9** *Assume  $(Y_t)$  is continuous and let  $a \in (y_0, \sup I]$ . Then both of the following two conditions imply (4.19)*

- (1)  $s(y_0) > 0$ ,  $s(Y_t) \rightarrow 0$  almost surely,  $Y_t < a$  for all  $t$  and  $s \circ \lambda(x) \geq 0$  for all  $x \in [y_0, a)$
- (2)  $s(\bar{Y}_\infty) = \Lambda(a) = \infty$

*Proof.* Under assumption (1) one has  $\bar{Y}_\infty < a$  and

$$\frac{s(\bar{Y}_t) - s(Y_t)}{s(\bar{Y}_t) - s \circ \lambda(\bar{Y}_t)} \geq \frac{s(\bar{Y}_t) - s(Y_t)}{s(\bar{Y}_t)} \rightarrow 1 \text{ almost surely.}$$

It follows that  $\bar{Y}_{T_\lambda} < a$  and  $Y_{T_\lambda} = \lambda(\bar{Y}_{T_\lambda})$ .

If condition (2) holds, then  $\int_0^\infty f(x)dx = \infty$  for the function

$$f(x) = 1_{\{0 \leq x < s(a) - s(y_0)\}} \frac{1}{x + s(y_0) - s \circ \lambda \circ s^{-1}(x + s(y_0))}.$$

Since  $X_t = s(\bar{Y}_t) - s(Y_t)$  is a non-negative continuous process of class  $(\Sigma)$  with decomposition  $X_t = (s(y_0) - s(Y_t)) + (s(\bar{Y}_t) - s(y_0))$ , it follows from Corollary 2.4 that  $\inf \{t : f(s(\bar{Y}_t) - s(y_0))X_t \geq 1\} < \infty$ . This implies  $T_\lambda < \infty$ , and (4.19) follows. □

We now are ready to extend our results on maximum drawdowns and maximum relative drawdowns of the last subsection.

**Proposition 4.10** *Assume  $(Y_t)$  is continuous with  $s(y_0) > 0$  and  $s(Y_t) \rightarrow 0$  almost surely. Let  $T < \infty$  be a stopping time and  $K$  an  $\mathcal{F}_T$ -measurable random variable such that  $0 \leq s(K) < s(Y_T)$ . Denote  $T_K = \inf \{t \geq T : Y_t = K\}$ . Then one has for all  $x \geq 0$ ,*

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t > x \mid \mathcal{F}_T \right] = 1_{\{\bar{Y}_{T-K} > x\}} + 1_{\{\bar{Y}_{T-K} \leq x\}} \frac{s(Y_T) - s(K)}{s(K+x) - s(K)}. \quad (4.24)$$

*If in addition,  $y_0 > 0$ , then*

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t > x \mid \mathcal{F}_T \right] = 1_{\{1 - K/\bar{Y}_T > x\}} + 1_{\{1 - K/\bar{Y}_T \leq x < 1\}} \frac{s(Y_T) - s(K)}{s(K/(1-x)) - s(K)}. \quad (4.25)$$

*Proof.* First assume that  $T = 0$  and  $K$  is a constant. Then  $(Y_t)$  with  $\lambda \equiv K$  and  $a = \sup I$  fulfills condition (1) of Proposition 4.9. Indeed,  $s(y_0) > 0$ ,  $s(Y_t) \rightarrow 0$  almost surely and  $s \circ \lambda(x) = s(K) \geq 0$  for all  $x \in [y_0, a)$  are part of the assumptions. To see that  $Y_t < a$  for all  $t$ , denote  $R = \inf \{t : s(Y_t) = 0\}$  and notice that  $M_t = s(Y_{t \wedge R})$  is a non-negative local martingale starting at  $s(y_0) > 0$  and converging to zero almost surely. Therefore, it follows from Doob's maximal identity that  $s(y_0)/s(\bar{Y}_R)$  is uniformly distributed on the interval  $(0, 1)$  (see Remark 4.3). In particular,  $s(a) = \infty$  and hence,  $Y_t < a$  for all  $t$ . It now follows from Proposition 4.9 that the conditions of Proposition 4.8 are fulfilled. Moreover,

$$\sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t = \bar{Y}_{T_K} - K \quad \text{and} \quad \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t = 1 - K/\bar{Y}_{T_K}.$$

So (4.24) and (4.25) can be deduced from formula (4.23). In the general case, the proposition follows by considering the process  $\tilde{Y}_t = Y_{T+t}$  in the filtration  $\tilde{\mathcal{F}}_t = \mathcal{F}_{T+t}$  and conditioning on  $\mathcal{F}_T$ .  $\square$

**Proposition 4.11** *Assume  $(Y_t)$  is continuous and  $s(\bar{Y}_\infty) = \infty$ . Let  $T < \infty$  be a stopping time and  $K$  a  $[0, \infty]$ -valued  $\mathcal{F}_T$ -measurable random variable such that  $\bar{Y}_T < K \leq \bar{Y}_\infty$ . Denote  $T_K = \inf \{t \geq T : Y_t = K\}$ . Then one has for all  $x \geq 0$ ,*

$$\mathbb{P}[Y_t \geq \lambda(\bar{Y}_t) \text{ for all } t \in [T, T_K] \cap \mathbb{R}_+ \mid \mathcal{F}_T] = \left( \frac{s(Y_T) - s \circ \lambda(\bar{Y}_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)} \right)^+ e^{\Lambda(\bar{Y}_T) - \Lambda(K)}. \quad (4.26)$$

*Proof.*  $X_t = s(\bar{Y}_t) - s(Y_t)$  is a non-negative process of class  $(\Sigma)$  with decomposition  $(s(y_0) - s(Y_t)) + (s(\bar{Y}_t) - s(y_0))$  and

$$f(x) = \frac{1}{x + s(y_0) - s \circ \lambda \circ s^{-1}(x + s(y_0))}$$

a non-negative Borel function from  $[0, s(\sup I) - s(y_0))$  to  $\mathbb{R}_+$ . Since  $Y_t \geq \lambda(\bar{Y}_t)$  is equivalent to  $f(s(\bar{Y}_t) - s(y_0))X_t \leq 1$  and  $T_K = \inf \{t \geq T : s(\bar{Y}_t) - s(y_0) = s(K) - s(y_0)\}$ , one obtains from Corollary 2.4 that

$$\begin{aligned} \mathbb{P}[Y_t \geq \lambda(\bar{Y}_t) \text{ for all } t \in [T, T_K] \cap \mathbb{R}_+ \mid \mathcal{F}_T] &= \left( \frac{s(Y_T) - s \circ \lambda(\bar{Y}_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)} \right)^+ \exp \left( - \int_{s(\bar{Y}_T) - s(y_0)}^{s(K) - s(y_0)} f(x) dx \right) \\ &= \left( \frac{s(Y_T) - s \circ \lambda(\bar{Y}_T)}{s(\bar{Y}_T) - s \circ \lambda(\bar{Y}_T)} \right)^+ \exp(\Lambda(\bar{Y}_T) - \Lambda(K)). \end{aligned}$$

$\square$

**Corollary 4.12** *If the assumptions of Proposition 4.11 hold, then*

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t \leq x \mid \mathcal{F}_T \right] = \left( \frac{s(Y_T) - s(\bar{Y}_T - x)}{s(\bar{Y}_T) - s(\bar{Y}_T - x)} \right)^+ \exp \left( - \int_{\bar{Y}_T}^K \frac{ds(y)}{s(y) - s(y-x)} \right) \quad (4.27)$$

for every constant  $x > 0$  such that  $y_0 - x \in \bar{I}$ . If in addition  $y_0 > 0$ , then

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t \leq x \mid \mathcal{F}_T \right] = \left( \frac{s(Y_T) - s([1-x]\bar{Y}_T)}{s(\bar{Y}_T) - s([1-x]\bar{Y}_T)} \right)^+ \exp \left( - \int_{\bar{Y}_T}^K \frac{ds(y)}{s(y) - s([1-x]y)} \right) \quad (4.28)$$

for each  $x > 0$  such that  $\inf \{(1-x)y : y \in [y_0, \sup I]\} \in \bar{I}$ .

*Proof.* Formula (4.27) follows from Proposition 4.11 applied to the function  $\lambda(y) = y - x$ . The condition  $y_0 - x \in \bar{I}$  ensures that  $\lambda([y_0, \infty) \cap I) \subset \bar{I}$ . Formula (4.28) is obtained from Proposition 4.11 applied to the function  $\lambda(y) = (1-x)y$ .  $\inf \{(1-x)y : y \in [y_0, \sup I]\} \in \bar{I}$  implies that  $\lambda([y_0, \infty) \cap I) \subset \bar{I}$ .  $\square$

## 5 Applications

We now discuss some applications in financial modelling and risk management. We consider a stochastic process modelling the evolution of a financial asset. In Subsection 5.1 below it is assumed to be a continuous local martingale  $(M_t)$ . In Subsection 5.2 it will be a solution of an SDE of the form  $dY_t = \mu(Y_t)dt + \sigma(Y_t)dB_t$ .

### 5.1 Pricing and hedging of options on running maxima

In standard mathematical finance it is usually assumed that there exists a risk-neutral measure under which discounted prices of tradable assets are non-negative local martingales. In the benchmark approach of Platen [21] (see also Platen and Heath [22]) prices are local martingales under the physical probability measure when expressed in terms of the benchmark portfolio. For our results to apply we also need the price  $(M_t)$  of an asset to be continuous and  $M_t \rightarrow 0$  almost surely as  $t \rightarrow \infty$ . So we here assume that  $(M_t)$  is a continuous non-negative local martingale with constant initial value  $m > 0$  such that  $M_t \rightarrow 0$  almost surely. We shall be interested in options that depend on the running maximum of  $(M_t)$  and are triggered by one of the following three stopping times:

1.  $T_c = \inf \{t : M_t = c\}$
2.  $T_c = \inf \{t : DD_t = c\}$
3.  $T_c = \inf \{t : rDD_t = c\}$

for a constant  $c \geq 0$ , where  $DD_t = \overline{M}_t - M_t$  and  $rDD_t = 1 - M_t/\overline{M}_t$ .

#### 5.1.1 Options with downfall triggers

Let us first consider an option with payoff of the form  $h(\overline{M}_{T_c})$  for a Borel function  $h : [m, \infty) \rightarrow \mathbb{R}$  and  $T_c = \inf \{t : M_t = c\}$  for some constant  $c \in [0, m)$ , that is, the option depends on the running maximum  $(\overline{M}_t)$  and is triggered the first time when  $(M_t)$  drops to  $c$ .  $T_c$  can be written as  $T_c = \inf \{t : M_t = \lambda(\overline{M}_t)\}$  for  $\lambda \equiv c$ . The functions  $\Lambda$  and  $h^\Lambda$  of Subsection 4.1 (see formulas (4.5) and (4.10)) corresponding to this particular  $\lambda$  take the form

$$\Lambda(x) = \log(x - c) - \log(m - c) \quad \text{and} \quad h^\Lambda(x) = (x - c) \int_x^\infty \frac{h(y)}{(y - c)^2} dy.$$

So it follows from Proposition 4.2 that if  $h$  satisfies the integrability condition  $\int_m^\infty |h(y)|/(y - c)^2 dy < \infty$ , one has

$$\mathbb{E} [h(\overline{M}_{T_c}) \mid \mathcal{F}_{t \wedge T_c}] = h^\Lambda(m) + \int_0^{t \wedge T_c} \frac{h^\Lambda(\overline{M}_u) - h^\Lambda(\overline{M}_u)}{\overline{M}_u - c} dM_u \quad (5.1)$$

$$= \frac{h(\overline{M}_{t \wedge T_c})DD_{t \wedge T_c} + h^\Lambda(\overline{M}_{t \wedge T_c})(M_{t \wedge T_c} - c)}{\overline{M}_{t \wedge T_c} - c}. \quad (5.2)$$

Formula (5.1) provides a hedging strategy and (5.2) the fair price of the option at time  $t \wedge T_c$ .

#### 5.1.2 Options with drawdown triggers

Now let the option payoff be given by  $h(\overline{M}_{T_c})$ , where  $T_c$  is the first time the drawdown  $(DD_t)$  hits some level  $c \in (0, m]$ . Then  $T_c$  can be written as  $T_c = \inf \{t : M_t = \lambda(\overline{M}_t)\}$  for the function  $\lambda(y) = y - c$ , and

the functions  $\Lambda$  and  $h^\lambda$  of Subsection 4.1 become

$$\Lambda(x) = (x - m)/c \quad \text{and} \quad h^\Lambda(x) = \frac{1}{c} e^{x/c} \int_x^\infty h(y) e^{-y/c} dy.$$

So it follows from Proposition 4.2 that if  $h$  satisfies  $\int_m^\infty |h(y)| e^{-y/c} dy < \infty$ , one has

$$\mathbb{E} [h(\overline{M}_{T_c}) | \mathcal{F}_{t \wedge T_c}] = h^\Lambda(m) + \frac{1}{c} \int_0^{t \wedge T_c} (h^\Lambda(\overline{M}_u) - h(\overline{M}_u)) dM_u \quad (5.3)$$

$$= h(\overline{M}_{t \wedge T_c}) \frac{DD_{t \wedge T_c}}{c} + h^\Lambda(\overline{M}_{t \wedge T_c}) \left( 1 - \frac{DD_{t \wedge T_c}}{c} \right). \quad (5.4)$$

Again, formula (5.3) gives the hedging strategy and (5.4) the fair price of the option.

### 5.1.3 Options with relative drawdown triggers

Consider an option with payoff  $h(\overline{M}_{T_c})$  for the stopping time  $T_c = \inf \{t : rDD_t = c\}$ , where  $c \in (0, 1]$ . Then  $T_c = \inf \{t : M_t = \lambda(\overline{M}_t)\}$  for  $\lambda(y) = (1 - c)y$ . The functions  $\Lambda$  and  $h^\Lambda$  of Subsection 4.1 then take the form

$$\Lambda(x) = \frac{1}{c} \log(x/m) \quad \text{and} \quad h^\Lambda(x) = \frac{1}{c} x^{1/c} \int_x^\infty h(y) y^{-(1+c)/c} dy,$$

and Proposition 4.2 gives for all functions  $h$  satisfying  $\int_m^\infty |h(y)| y^{-(1+c)/c} dy < \infty$ ,

$$\begin{aligned} \mathbb{E} [h(\overline{M}_{T_c}) | \mathcal{F}_{t \wedge T_c}] &= h^\Lambda(m) + \int_0^{t \wedge T_c} \frac{h^\Lambda(\overline{M}_u) - h(\overline{M}_u)}{c \overline{M}_u} dM_u \\ &= h(\overline{M}_{t \wedge T_c}) \frac{rDD_{t \wedge T_c}}{c} + h^\Lambda(\overline{M}_{t \wedge T_c}) \left( 1 - \frac{rDD_{t \wedge T_c}}{c} \right), \end{aligned}$$

showing how to hedge and price the option.

## 5.2 Risk management

Risk managers and regulators are typically interested in the distribution of prices under the physical probability measure. In standard mathematical finance they are assumed to follow semimartingales. In the benchmark approach of Platen [21] they are local martingales. Let us here consider a price process  $(Y_t)$  taking values in an interval  $I \subset \mathbb{R}$  and satisfying an SDE of the form  $dY_t = \mu(Y_t)dt + \sigma(Y_t)dB_t$ ,  $Y_0 = y_0 \in I$  for a Brownian motion  $(B_t)$  such that

$$\gamma(x) = 2 \int_{y_0}^x \frac{\mu(y)}{\sigma^2(y)} dy \quad \text{and} \quad \int_{y_0}^x e^{-\gamma(y)} dy \quad \text{are finite for all } x \in I.$$

Choose constants  $c \in \mathbb{R}$ ,  $d > 0$  and set  $s(x) = c + d \int_{y_0}^x e^{-\gamma(y)} dy$ . Then the scaled process  $s(Y_t)$  is a local martingale. Let  $\lambda : [y_0, \infty) \cap I \rightarrow \bar{I}$  be a Borel function satisfying the assumptions of Proposition 4.8 for some  $a \leq \infty$  and denote

$$T_\lambda = \inf \{t : Y_t = \lambda(\overline{Y}_t)\} \quad \text{and} \quad g_\lambda = \sup \{t \leq T_\lambda : Y_t = \overline{Y}_t\}.$$

The function  $\Lambda$  defined in (4.18) then becomes

$$\Lambda(x) = \int_{y_0}^x \frac{e^{-\gamma(y)} dy}{\int_{\lambda(y)}^y e^{-\gamma(z)} dz},$$

and one obtains from Proposition 4.8 that for all stopping times  $T \leq T_\lambda$ ,

$$\mathbb{P}[g_\lambda \leq T \mid \mathcal{F}_T] = \frac{\int_{Y_T}^{\bar{Y}_T} e^{-\gamma(y)} dy}{\int_{\lambda(\bar{Y}_T)}^{\bar{Y}_T} e^{-\gamma(y)} dy}$$

and

$$\mathbb{P}[\bar{Y}_{T_\lambda} > x \mid \mathcal{F}_T] = 1_{\{\bar{Y}_T > x\}} + 1_{\{\bar{Y}_T \leq x\}} \frac{\int_{\lambda(\bar{Y}_T)}^{Y_T} e^{-\gamma(y)} dy}{\int_{\lambda(\bar{Y}_T)}^{\bar{Y}_T} e^{-\gamma(y)} dy} \exp\left(-\int_{\bar{Y}_T}^x \frac{e^{-\gamma(y)} dy}{\int_{\lambda(y)}^y e^{-\gamma(z)} dz}\right) \quad \text{for } x \geq y_0. \quad (5.5)$$

In the special case  $\lambda \equiv c$ , (5.5) reduces to

$$\mathbb{P}[\bar{Y}_{T_\lambda} > x \mid \mathcal{F}_T] = 1_{\{\bar{Y}_T > x\}} + 1_{\{\bar{Y}_T \leq x\}} \frac{\int_c^{\bar{Y}_T} e^{-\gamma(y)} dy}{\int_c^x e^{-\gamma(y)} dy} \quad \text{for } x \geq y_0,$$

and for  $\lambda(y) = y - c$  it becomes

$$\mathbb{P}[\bar{Y}_{T_\lambda} > x \mid \mathcal{F}_T] = 1_{\{\bar{Y}_T > x\}} + 1_{\{\bar{Y}_T \leq x\}} \frac{\int_{Y_T - c}^{Y_T} e^{-\gamma(y)} dy}{\int_{\bar{Y}_T - c}^{\bar{Y}_T} e^{-\gamma(y)} dy} \exp\left(-\int_{\bar{Y}_T}^x \frac{e^{-\gamma(y)} dy}{\int_{y-c}^y e^{-\gamma(z)} dz}\right) \quad \text{for } x \geq y_0.$$

For  $T = 0$ , this gives

$$\mathbb{P}[\bar{Y}_{T_\lambda} > x] = \exp\left(-\int_{y_0}^x \frac{e^{-\gamma(y)} dy}{\int_{y-c}^y e^{-\gamma(z)} dz}\right) \quad \text{for } x \geq y_0,$$

which (in the case  $y_0 = 0$ ) is formula (3) of Lehoczy [12].

If  $s(y_0) > 0$ ,  $s(Y_t) \rightarrow 0$  almost surely and there exists an  $\mathcal{F}_T$ -measurable random variable  $K$  such that  $0 \leq s(K) < s(Y_T)$ , denote  $T_K = \inf\{t \geq T : Y_t = K\}$ . Then it follows from Proposition 4.10 that for all  $x \geq 0$ ,

$$\mathbb{P}\left[\sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t > x \mid \mathcal{F}_T\right] = 1_{\{\bar{Y}_{T-K} > x\}} + 1_{\{\bar{Y}_{T-K} \leq x\}} \frac{\int_K^{Y_T} e^{-\gamma(y)} dy}{\int_K^{K+x} e^{-\gamma(y)} dy}.$$

If in addition,  $y_0 > 0$ , then

$$\mathbb{P}\left[\sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t > x \mid \mathcal{F}_T\right] = 1_{\{1-K/\bar{Y}_T > x\}} + 1_{\{1-K/\bar{Y}_T \leq x < 1\}} \frac{\int_K^{Y_T} e^{-\gamma(y)} dy}{\int_K^{K/(1-x)} e^{-\gamma(y)} dy}.$$



On the other hand, if

$$\int_{y_0}^{\bar{Y}_\infty} e^{-\gamma(y)} dy = \infty$$

and  $K$  is a  $[0, \infty]$ -valued  $\mathcal{F}_T$ -measurable random variable such that  $\bar{Y}_T < K \leq \bar{Y}_\infty$  we denote  $T_K = \inf \{t \geq T : Y_t = K\}$  and obtain from Corollary 4.12 that for all  $x \geq 0$ ,

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} DD_t \leq x \mid \mathcal{F}_T \right] = \left( \frac{\int_{\bar{Y}_T - x}^{Y_T} e^{-\gamma(y)} dy}{\int_{\bar{Y}_T - x}^{\bar{Y}_T} e^{-\gamma(y)} dy} \right)^+ \exp \left( - \int_{\bar{Y}_T}^K \frac{e^{-\gamma(y)} dy}{\int_{y-x}^y e^{-\gamma(z)} dz} \right)$$

for every constant  $x > 0$  such that  $y_0 - x \in \bar{I}$ . If in addition  $y_0 > 0$ , then

$$\mathbb{P} \left[ \sup_{t \in [T, T_K] \cap \mathbb{R}_+} rDD_t \leq x \mid \mathcal{F}_T \right] = \left( \frac{\int_{(1-x)\bar{Y}_T}^{Y_T} e^{-\gamma(y)} dy}{\int_{(1-x)\bar{Y}_T}^{\bar{Y}_T} e^{-\gamma(y)} dy} \right)^+ \exp \left( - \int_{\bar{Y}_T}^K \frac{e^{-\gamma(y)} dy}{\int_{(1-x)y}^y e^{-\gamma(z)} dz} \right)$$

for each  $x > 0$  such that  $\inf \{(1-x)y : y \in [y_0, \sup I]\} \in \bar{I}$ .

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